Tracking and Detecting moving weak Targets

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ABSTRACT

Detecting and tracking of moving weak targets is a complicated dynamic state estimation problem whose difficulty is increased in case of high clutter conditions or low signal to noise ratio (SNR). In this case, the track-before-detect filter (TBDF) that uses unthresholded measurements considers as an effective method for detecting and tracking a single target under low SNR conditions. In this paper, a particle filter based track-before-detect (PF-TBD) method is proposed to address the problem of detection and tracking with unthresholded data and a binary variable of the existence of the target for two motion models. Simulation results using image measurements based on TBD scenarios are also presented to demonstrate the capability of the proposed approach.

1. Introduction

The classical approach to target tracking is based on target measurements (position, range rate, and so forth) that are extracted by thresholding the output of a signal processing unit of a surveillance sensor [1]. The primary role of thresholding is to reduce the data flow and thus simplify tracking. For a target of a certain signal-to-noise ratio (SNR), the choice of the detection threshold determines the probability of target detection and the density of false alarms. The false alarm rate, on the other hand, affects the complexity of the data association problem in the tracking system. In general, higher densities of false alarms require more sophisticated data association algorithms.

The undesirable effect of thresholding the sensor data, however, is that in restricting the data flow, it also throws away potentially useful information. For high SNR targets this loss of information is of little concern because one can achieve good probability of detection with a small false alarm rate. Recent developments of stealthy military aircraft and cruise missiles have emphasized the need to detect and track low SNR targets. For these dim (stealthy) targets, there is a considerable advantage in using the unthresholded data for simultaneous detection and track initiation [2], [3]. Depending on the type of sensor in use, the unthresholded data can be a sequence of range-Doppler maps, bearing-frequency distributions.

The concept of simultaneous detection and tracking using unthresholded data is known in literatures as track-before-detect (TBD) approach. Typically TBD is implemented as a batch algorithm using the Hough transform [4], dynamic programming [2] [3] or maximum likelihood estimation [5].

TBD algorithms based on the Hough transformation, dynamic programming or maximum likelihood methods are generally computationally intensive [6]. With recent advancement in Sequential Monte Carlo techniques, TBD algorithms implemented using PF are now computationally feasible [7] [8].

In this paper we also develop a recursive Bayesian TBD estimator; however, our formulation and implementation are based on the particle filter [7] [9]. The PF based TBD incorporates unthresholded data and a binary target existence variable into the target state estimation process. The presence and absence of target are explicitly modelled [10] [11]. This concept, allows us to calculate the probability of existence of the target directly from the filter.

The paper is organized as follows: section 2 the system dynamics and measurement model, are introduced for the TBD application. In section 3 formulates the TBD approach as a nonlinear filtering problem and describes the conceptual recursive Bayesian solution. The implementation of this solution using a particle filter is presented in section 4. In section 5 collects our simulations and results. Finally, we report our conclusions and direction for future research in section 6.
2. Formulation Problem

2.1. Dynamic Model

We assume that want to track a target moving in a 2-D plane with an unknown state vector $s_k$ at time step $k$. We consider the state model given by:

$$s_{k+1} = Fs_k + v_k$$

(1)

Where $F$ is the state transition matrix, assuming constant velocity motion and coordinate turn motion respectively, $k$ is the discrete-time index, $v_k$ is the process noise sequence, and $s_k$ is the state vector defined as:

$$s_k = [x_k \ y_k \ \dot{x}_k \ \dot{y}_k \ \Delta x_k \ \Delta y_k]$$

(2)

Here $(x_k, y_k), (\dot{x}_k, \dot{y}_k)$ and $I_k$ denote the position, velocity, and the intensity of the target, respectively.

2.2. Transition Matrix

A target can be present or absent from the surveillance region at a discrete-time $k$. Target presence variable $E_k$ is modelled by a two-state Markov chain, that is $E_k = \{0,1\}$. Here 0 denotes the event that a target is not present, while 1 denotes the opposite. Furthermore, we assume that transitional probabilities of target “birth” $P_b$ and “death” $P_d$, defined as:

$$P_b = P[E_k = 1|E_{k-1} = 0]$$

(3)

$$P_d = P[E_k = 0|E_{k-1} = 1]$$

(4)

Are known, the other two transitional probabilities of this Markov chain, the probability of staying alive $1 - P_d$ and the probability of remaining absent $1 - P_b$ respectively, are given by:

$$1 - P_d = P[E_k = 1|E_{k-1} = 1]$$

(5)

$$1 - P_b = P[E_k = 0|E_{k-1} = 0]$$

(6)

The corresponding transition matrix for the Markov process is:

$$\Pi = \begin{bmatrix} 1-P_b & P_b \\ P_d & 1-P_d \end{bmatrix}$$

(7)

2.3. Sensor Model

The sensor provides a sequence of two-dimensional images (frames) of the surveillance region, each image consisting of $(n \times m)$ resolution cells (pixels). A resolution cell corresponds to a rectangular region of dimensions $\Delta x \times \Delta y$, so that the center of each cell $(i,j)$ is defined to be at $(i\Delta x \times j\Delta y)$ for $i = 1, \ldots, n$ and $j = 1, \ldots, m$.

At each resolution cell $(i,j)$ the measured intensity is denoted as $z_k^{(i,j)}$ and modeled as:

$$z_k^{(i,j)} = \begin{cases} h_k^{(i,j)}(s_k) + \omega_k^{(i,j)} & \text{if target present} \\ \omega_k^{(i,j)} & \text{if target absent} \end{cases}$$

(8)

Where $h_k^{(i,j)}(s_k)$ is the target contribution to intensity level in the resolution cell $(i,j)$ and $\omega_k^{(i,j)}$ is measurement noise in the resolution cell $(i,j)$, assumed to be independent from pixel to pixel and from frame to frame. Thus for a point target of intensity $I_k$ at position$(x_k, y_k)$, the contribution to pixel$(i,j)$ is approximated as:

$$h_k^{(i,j)}(s_k) = \frac{\Delta x \Delta y I_k}{2\pi \sigma^2} \exp \left\{ -\frac{(\Delta x x_k)^2 + (\Delta y y_k)^2}{2\sigma^2} \right\}$$

(9)

Where $\Sigma$ is the amount of blurring introduced by the sensor. The complete measurements recorded at time $k$ a $n \times m$ matrix denoted as:

$$z_k = \{z_k^{(i,j)}\}_{i=1}^{n}$$

(10)

While the set of complete measurements collected up to time $k$ is denoted as usual: $Z_k = \{z_i, i = 1, \ldots, k\}$.

3. Bayesian Solution to TBD Filtering

The formal recursive Bayesian solution can be presented as a two-step procedure, consisting of prediction and update.

3.1. Prediction

The predicted target state can be written in terms of the target state and existence at the previous time, giving:

$$p(s_k, E_k = 1|Z_{k-1}) = (1 - P_d) + P_d \{ p(s_{k-1}|s_{k-1}, E_{k-1} = 1) \times p(s_{k-1} = 0|Z_{k-1}) ds_{k-1} +$$

$$P_b \{ p_b(s_k)p(s_{k-1}, E_{k-1} = 0|Z_{k-1}) ds_{k-1} \}$$

(11)

The pdf $p_b(s_k)$ denotes the initial target density on its appearance.

3.2. Update

The update equation in the Bayesian framework is given by:

$$p(z_k|s_k, E_k = 1|Z_{k-1}) = \frac{p(s_k, E_k = 1|Z_{k-1})}{p(z_k|Z_{k-1})}$$

(12)

Where prediction density $p(s_k, E_k = 1|Z_{k-1})$ is given by (11) and $p(z_k|s_k, E_k)$ is the likelihood function given by:

$$p(z_k|s_k, E_k) =$$

$$\begin{cases} p(s_k|s_{k-1}, E_k = 1)p(s_{k-1}, E_k = 1|Z_{k-1}) & \text{for } E_k = 1 \\ p(s_k|s_{k-1}, E_k = 0)p(s_{k-1}, E_k = 0|Z_{k-1}) & \text{for } E_k = 0 \end{cases}$$

(13)

Here $p_N(z_k^{i,j})$ is the probability density function of background noise in pixel $(i,j)$, while $p_{S+N}(z_k^{i,j}|s_k)$ is the likelihood of target signal plus noise in pixel $(i,j)$ given that the target is in state $s_k$. This two probability density function can be further expressed as:

$$p_N(z_k^{i,j}) = N(z_k^{i,j}, 0, \sigma^2)$$

(14)
\[
p_{s+n} \left( z_{k}^{(i,j)} \middle| s_{k} \right) = N \left( z_{k}^{(i,j)}, h_{k}^{(i,j)}, \sigma^2 \right) \tag{15}
\]

Since the target (if present) will affect only the pixels in the vicinity of its location \((x_k, y_k)\), the expression for \(p(z_k | s_k, E_k = 1)\) can be approximated as follows:

\[
p(z_k | s_k, E_k = 1) \approx \prod_{i \in \mathcal{C}(s_k)} \prod_{j \in \mathcal{C}(s_k)} p_{s+n} \left( z_{k}^{(i,j)} \middle| s_{k} \right) \cdot \prod_{\forall i \in \mathcal{C}(s_k)} \prod_{\forall j \in \mathcal{C}(s_k)} p_{n} \left( z_{k}^{(i,j)} \right) \tag{16}
\]

Where \(\mathcal{C}(s_k)\) and \(\mathcal{C}(s_k)\) are the sets of subscripts \(i\) and \(j\), respectively, corresponding to pixels affected by the target.

4. A Particle Filter for Track Before Detect (PF-TBD)

The recursive Bayesian solution of the track problem described in the previous section can be implemented using a particle filter [7] [9] [12] [13] has some similarities to the MMPF [14]. In this case, we introduce the augmented state vector to include the existence variable, \(y_k = [s_k^T E_k^T] \). Let us denote a random measure that characterizes the posterior probability density function at \(k - 1\), namely \(p(y_k | z_{k-1})\), by \(\omega_{k-1}^n, \omega_{k-1}^n\). As usual, \(N\) is the number of particles, while \(y_k^n\) consists of \(s_k^n\) and \(E_k^n\). The pseudocode of a single cycle of the PF developed for the TBD problem is presented in Table 1. The next step is the prediction of particle target states; this is done, however, only for those particles that are characterized by \(E_k^n = 1\). For remaining particles (with \(E_k^n = 0\)), the target state components are undefined. There are two possible cases here:

4.1. Newborne Particles

This group of predicted particles is characterized by the transition from \(E_{k-1}^n = 0\) to \(E_k^n = 1\). The target state particles are uniformly drawn at time step \(k\) based on some a priori information on the minimum and maximum possible values on the target state.

4.2. Existing Particles

This group of particles that continues to stay “alive”, with \(E_{k-1}^n = 1\) to \(E_k^n = 1\). The state transition model in equation (1) is used to update the target state particles.

The importance weights are computed next. For this purpose we need to introduce the likelihood ratio in pixel \((i, j)\) for a target in state \(s^n_k\), defined as:

\[
\ell \left( z_{k}^{(i,j)} \middle| s_{k}^{n} \right) = \frac{p_{s+n} \left( z_{k}^{(i,j)} \middle| s_{k}^{n} \right)}{p_{n} \left( z_{k}^{(i,j)} \right)} \tag{17}
\]

\[
\ell \left( z_{k}^{(i,j)} \middle| s_{k}^{n} \right) = \exp \left( \frac{-h_{k}^{(i,j)}(s_{k}^{n} - 2z_{k}^{(i,j)})}{2\sigma^2} \right) \tag{18}
\]

Where \(h_{k}^{(i,j)}\) was defined in (9). Equation (18) follows from (14), (15), and (11). The importance weights (up normalizing constant) are now given by [7]:

\[
\begin{cases}
\tilde{\omega}_{k}^{n} = \prod_{\forall i \in \mathcal{C}(s_{k}^{n})} \prod_{\forall j \in \mathcal{C}(s_{k}^{n})} \ell \left( z_{k}^{(i,j)} \middle| s_{k}^{n} \right) & \text{if } E_{k}^{n} = 1 \\
1 & \text{if } E_{k}^{n} = 0
\end{cases} \tag{19}
\]

The PF for track-before-detect performs target detection using the estimate of the posterior probability of target existence. This estimate is computed as:

\[
\hat{P}_{k} = \frac{\sum_{n=1}^{N} E_{k}^{n}}{N} \tag{20}
\]

And satisfies \(0 < \hat{P}_{k} < 1\). Target presence is declared if \(\hat{P}_{k}\) is above a certain threshold value. This declaration can then trigger the initialization of a track based on the estimated target state.

\[
\hat{s}_{k | k} = \frac{\sum_{n=1}^{N} E_{k}^{n} s_{k}^{n}}{\sum_{n=1}^{N} E_{k}^{n}} \tag{21}
\]

5. Simulation and Results

In the simulation we used two scenarios for target motion and random walk model is adopted for target intensity.

5.1. Scenario 1

The first model is a nearly constant velocity model is used. The dynamic model [15] for the target can be described by (1).

Where

\[
F = \text{diag}[F_1, F_1], \quad Q = \text{diag}[Q_1, Q_1, q_2T]
\]

and

\[
F_1 = \begin{bmatrix} 1 & T \\ 0 & 1 \end{bmatrix}, \quad Q_1 = q_1 \begin{bmatrix} T^3/3 & T^2/2 \\ T^2/2 & T \end{bmatrix}
\]

Where \(q_1\) and \(q_2\) denote the level of process noise in target motion and intensity, respectively. A sequence of 30 frames of data has been generated with the following parameters:

\[
\Delta x = \Delta y = 1, n = m = 20, T = 1, \sigma = 3, \Sigma = 0.7
\]

The target is absent from frame 1 to frame 5 to be introduced in frame 6 with the initial intensity \(I_0\). The initial state is [4.2 0.45 7.2 0.25 0.05] [6] [16] [17].
The simulations are conducted under an initial intensity $I_0 = 9, 13$ and 25, which corresponds to an SNR of 3.18, 6.71, and 12 dB, respectively, according to the calculation equation $SNR = 10\log\left(\frac{I_{peak}}{\sigma^2}\right)$. The target exists until frame 24 and is again absent in frames 25, 26, ..., 30.

Figure 1 (a) and (b) show the measurement frame at time step 20 for 6.71 dB and 12 dB peak respectively.

![Figure 1](www.astesj.com)

Figure 2. True and Estimated Target Existence Probability for SNR=6.71 dB

Figure (3) displays the true target path against the track, produced by the filter. Note how the target trajectory deviates slightly from the straight line due to process noise. The PF-TBD tracks the target with a small positional error.

![Figure 3](www.astesj.com)

Figure 3. True and Estimated Target Trajectory for SNR=6.71 dB

5.2. Scenario 2

The second model is a Coordinate turn model is used [15]. The dynamic model for the target can be described by (1).

$$F = \begin{bmatrix} 1 & F_1 & 0 & F_2 & 0 \\ 0 & F_3 & 0 & -F_4 & 0 \\ 0 & -F_2 & 1 & F_3 & 0 \\ 0 & F_4 & 0 & F_3 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{bmatrix},$$

$$Q = \begin{bmatrix} Q_{11} & Q_{12} & 0 & 0 \\ Q_{21} & Q_{22} & -Q_3 & 0 \\ 0 & -Q_3 & Q_{12} & 0 \\ 0 & Q_3 & 0 & Q_{12} \\ 0 & 0 & 0 & 0 \end{bmatrix}.$$

And $F_1 = \frac{\sin(\Psi T)}{\psi}, F_2 = \frac{\cos(\Psi T) + 1}{\psi}, F_3 = \cos(\Psi T), F_4 = \frac{\sin(\Psi T)}{\psi}, Q_1 = \frac{2 (\Psi T - \sin(\Psi T))}{\psi^3}, Q_2 = \frac{(1 - \cos(\Psi T)) \psi}{\psi^2}, Q_3 = c T, Q_4 = q_1 T, Q_5 = q_2 T, \psi = 6$ is a constant angular rate.

Figure 5 (a) and (b) show the measurement frame at time step 20 for 6.71 dB and 12 dB peak respectively.

![Figure 5](www.astesj.com)

Figure 5. Measurements Frame (20): (a) for 6.71 dB Peak SNR, (b) for 12 dB Peak SNR for CT model.

In figure (6) the probability of presence is shown for a SNR of 6.71 dB. Existence probability probability is still increase above frame 7 until frame 17 and still stable until frame 25. Therefore, it drops rapidly following the target disappears from the monitoring region after frame 25.

Figure (7) shows the true and estimated target trajectories for coordinate turn model, the estimated trajectory is very close to the true trajectory.

Figure (8) shows the position RMSE for three different peak SNR conditions (3.18 dB, 6.71 dB, 12 dB). The position error is lower in 6.71 dB than 3.18 dB. As it can be seen, the PF-TBD was able to closely track the target even under low SNR.
Fig. 6. True and Estimated Target Trajectory for SNR=6.71 dB

Fig. 7. True and Estimated Target CT Trajectory for SNR=6.71 dB

Fig. 8. Position RMSE for the PF-TBD for different peak SNR

6. Conclusion

In this paper, to manipulate moving weak targets, the PF-TBD algorithm is proposed for two dynamics models (CV and CT). The major advantage of the track-before-detect approach based on target existence variable and as a result, the developed particle filter can detect and track low SNR maneuvering target. The results from the simulation show that the PF-TBD algorithm has a successfully detection and tracking performance, both for constant velocity and coordinate turn models of moving targets, under severe conditions such as high noise or low SNR. Therefore, further work will mainly concentrate on how to detect and track multiple targets in high noise and high clutter.

References